

Lecture 2

Probability
Space
Formalism

$(\Omega, \mathcal{F}, \mathbb{P})$ — events (a σ -algebra on Ω)
— Probability measure (axioms)
— set (of samples)

A I. $A \in \mathcal{F} \Rightarrow P(A) \geq 0$ ← negativity

A II. $P(\Omega) = 1$ ← $P(\text{anything happening}) = 1$

A III. A_1, A_2, \dots disjoint events ← additivity
 $\Rightarrow P(\bigcup_{i=1}^{\infty} A_i) := \sum_{i=1}^{\infty} P(A_i)$

Properties (derived last lecture):

- $A \in \mathcal{F} \Rightarrow P(A^c) = 1 - P(A)$
- $A \subset B \Rightarrow P(A) \leq P(B)$
- $P(A \cup B) = P(A) + P(B) - P(A \cap B)$ } inclusion-exclusion
- $P(B) = \sum_{i=1}^{\infty} P(B \cap A_i)$ for $A_1, A_2, \dots \in \mathcal{F}$ a partition of Ω
↳ (MECE) ↳ total probability rule

Today: Conditional Probability, Independence, Random Vars

Def: If $B \in \mathcal{F}$, $P(B) > 0$, then we define

conditional probability := $\frac{P(A \cap B)}{P(B)}$ $A \in \mathcal{F}$

↳ intuition: $P(A|B) =$ "probability A occurs given we know that B occurred"

↳ more formally: $P(\cdot|B)$ gives a restriction of our model $(\Omega, \mathcal{F}, \mathbb{P})$ to those samples in B

↳ i.e., this triple $(B, \mathcal{F}|_B, P(\cdot|B))$ is a probability space itself, where

$$\mathcal{F}|_B = \{A \cap B : A \in \mathcal{F}\}$$

Example (of Cond. Prob):

If $A_1, A_2, \dots \in \mathcal{F}$, $P(A_i) > 0$, (A_i) partition Ω ,

↳ we can rewrite the law of total probability in terms of conditional probabilities:

$$P(B) = \sum_i P(B \cap A_i) = \sum_i P(B | A_i) P(A_i)$$

Bayes Rule

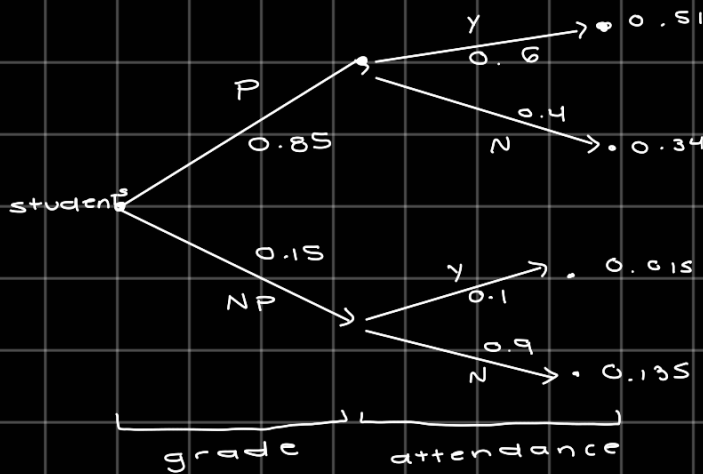
• importance: sometimes it's easy to express $P(B | A)$ but we're interested in $P(A | B)$

this is the task of inference: how do I get my state from an observation?

$$P(A | B) = \frac{P(B \cap A)}{P(B)} = \frac{P(A) P(B | A)}{P(B)}$$

examples

① Suppose we have the following model:



Question:

$$\frac{P(NP | N)}{P(NP | Y)} = ?$$

$$\frac{P(NP | N)}{P(NP | Y)} = \frac{P(NP \cap N) P(Y)}{P(NP \cap Y) P(N)} = \frac{0.135 \cdot 0.51}{0.015 \cdot 0.48} = 9.9$$

achievement gap

"How much more likely are students to NP given their attendance?"
9.9x

② Dice Rolling: we roll 2 dice & the sum is 10.

What is the probability that roll 1 was = 4?

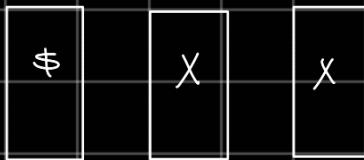
$$B = \{ \text{sum of rolls} = 10 \}$$

$$A = \{ \text{First roll} = 4 \}$$

$$P(A | B) = \frac{P(A \cap B)}{P(B)} = \frac{P(\{(4, 6)\})}{P(\{(4, 6), (5, 5), (6, 4)\})} = \frac{1/36}{3/36} = \frac{1}{3}$$

sample space: all possible tuples of rolls → 36 elements

© Monty hall problem



1 → pick door at random $\frac{1}{3}$ probability

2 → goat!

3 → switch

• conditioning ^{also} gives us a way to ^{usefully} decompose intersections of events

Consider: event $A_1 \cap A_2 \cap \dots \cap A_n$

$$P\left(\bigcap_{i=1}^n A_i\right) = P\left(A_1 \mid \bigcap_{i=2}^n A_i\right) P\left(\bigcap_{i=2}^n A_i\right)$$

$$= P(A_1) P\left(\bigcap_{i=2}^n A_i \mid A_1\right)$$

$$= P(A_1) P(A_2 \mid A_1) P(A_3 \mid A_1 \cap A_2) \dots P(A_n \mid \bigcap_{i=1}^{n-1} A_i)$$

Birthday Paradox

• given n people in a room, what is probability at least 2 share a birthday?

$A_i = \{ \text{person } i \text{ doesn't share a birthday w/ people}$

$j = 1 \dots i-1 \}$ $\leadsto 365 - (i-1)$ options for bdays

$$P\left(A_i \mid \bigcap_{j=1}^{i-1} A_j\right) = 1 - \frac{(i-1)}{365} = \frac{365 - (i-1)}{365}$$

$$P(\text{No shared b-days}) = P\left(\bigcap_{i=1}^n A_i\right) = \prod_{i=1}^n P\left(A_i \mid \bigcap_{j=1}^{i-1} A_j\right)$$

$$= \prod_{i=1}^n \left(1 - \frac{(i-1)}{365}\right)$$

$$\leq e^{-\sum_{i=1}^n \frac{i-1}{365}} = e^{-\binom{n}{2}/365}$$

Recall $1 - x \leq e^{-x}$

* as an exercise for yourself, find the probability space
↳ in this problem, it was implicit

$$\Rightarrow P(\text{shared bday}) \geq 1 - e^{-\binom{n}{2}/365}$$

$$\leadsto n = 23 \rightarrow P(\text{shared}) = 0.5$$

$$n = 261 \rightarrow P(\text{shared}) = 0.9999 \dots$$

Independence

- def: events A, B are independent if $P(A \cap B) = P(A)P(B)$

↳ misconception: that disjoint sets are independent

- NOTE: in special case where

$$P(A) > 0 \quad \& \quad A, B \text{ independent} \quad \Leftrightarrow \quad P(B|A) = P(B)$$

- in general:

collection $A_1, A_2, \dots \in \mathcal{F}$ are independent events if

$$P\left(\bigcap_{i \in S} A_i\right) = \prod_{i \in S} P(A_i) \quad \forall \text{ finite sets of indices } S.$$

subtle

↳ this is why we don't talk about *only* many iid coin flips ... how do you even conceptualize that?

- If $A_1, A_2, \dots \in \mathcal{F}$ are independent, then B_1, B_2, \dots are independent where each $B_i = A_i$ or A_i^c . Are the B_i 's independent?

↳ idea: if you take the complements of events, the complements are also independent

Proof:

$$\bigcap_{i=2}^n A_i = \left(\bigcap_{i=1}^n A_i\right) \cup \left(A_i^c \cap \bigcap_{i=2}^n A_i\right)$$

since these are independent, can rewrite as sum of probabilities:

def in appearance \rightarrow

$$\prod_{i=2}^n P(A_i) = P\left(\bigcap_{i=2}^n A_i\right) = P\left(\left(\bigcap_{i=1}^n A_i\right) \cup \left(A_i^c \cap \bigcap_{i=1}^n A_i\right)\right)$$
$$\rightarrow = P\left(\bigcap_{i=1}^n A_i\right) + P\left(A_i^c \cap \bigcap_{i=2}^n A_i\right)$$

↳ collecting terms & rewriting

$$(1 - P(A_1)) \prod_{i=2}^n P(A_i) = P\left(A_1^c \cap \bigcap_{i=2}^n A_i\right)$$

$P(A_1^c)$

$\Rightarrow A_1^c, \dots, A_n$ are independent

↳ by def of independence $\Rightarrow B_1, \dots, B_n$ independent

can't do induction past ∞ , must do it for finite sets

Conditional independence

- def: if $A, B, C \in \mathcal{F}$ are such that $P(C) > 0$

and $P(A \cap B \cap C) = P(A|C)P(B|C)$, then A, B are said to be conditionally independent

given C .

- ex: 2 coins with bias $p \neq q$ with probability
↳ pick coin at random (w.p. $1/2$), flip twice

$$H_i = \{\text{Flip } i = H\}$$

using law of total probability w/ conditional probabilities

$$P(H_i) = \frac{p+q}{2}$$

$$P(H_1 \cap H_2) = \frac{p^2 + q^2}{2} \neq P(H_1)P(H_2) = \left(\frac{p+q}{2}\right)^2$$

$$C = \{\text{pick coin } p\}$$

↳ when H_1, H_2 conditionally independent given C .